

Product, Ratio and Single Moments of Lower Record Values of Inverse Weibull Distribution

M. Aleem

Abstract:

In this paper the product, ratio and single moments of the Lower Record values are obtained from Inverse Weibull distribution.

Key Words:

Moments, Recurrence Relations, Lower Record values, Inverse Weibull distribution.

1. Introduction:

A random variable X has an Inverse Weibull distribution with pdf given by

$$f(x) = \frac{m}{\theta x^{m+1}} \exp\left(-\frac{1}{\theta x^m}\right) \quad (1.1)$$

where $x > 0, (\theta, m) > 0$

And the corresponding cdf is given by

$$F(x) = \exp\left(-\frac{1}{\theta x^m}\right) \quad (1.2)$$

where $x > 0, (\theta, m) > 0$

If we put $m=1$, it reduces to the Inverse Exponential distribution. If we put $m=2$, it reduces to the Inverse Rayleigh distribution. Some work has been done on Inverse Rayleigh distribution by Voda (1972), Ghanaph (1993), Mukargee & Saren (1984) and Mukarjee & Mait (1996). For distributional properties of Record values of inverse weibull distribution see Aleem and Pasha (2003), Aleem(2004). Other references are Ahsanullah (1995), Ahsanullah and Novzorov (2001).

Let $v(x) = \frac{1}{\theta x^m}$ and from (1.1) (1.2), we have

$$F(x) = \frac{1}{v^*(x)} \cdot f(x)$$

2. Product Moments

The lower record values are respected by $X_{L(1)}, X_{L(2)}, \dots, X_{L(n)}$

The joint pdf of $X_{L(r)}$ and $X_{L(s)}$ ($s > r$) is

$$f_{(r),(s)}(x,y) = C_{r,s} [H(x)]^{r-1} [H(y) - H(x)]^{s-r-1} h(x) \cdot f(y) \quad (2.1)$$

$$\text{where } C_{r,s} = \frac{1}{(r-1)!(s-r-1)!} \quad -\infty < y < x < \infty$$

$$\text{and } H(x) = -\text{Ln}F(x) \quad 0 < F(x) < 1$$

$$h(x) = -\frac{d}{dx} H(x)$$

If g is a Borel measurable function from \mathbb{R}^2 to \mathbb{R} , then

$$\begin{aligned} E\{g(X_{L(r)}, X_{L(s)})\} &= \\ C_{r,s} \int_0^{\infty} \int_{0 < y < x < \infty} g(x,y) [H(x)]^{r-1} [H(y) - H(x)]^{s-r-1} & \\ h(x) f(y) dx dy & \quad (2.2) \end{aligned}$$

Theorem 2.1 :

For the distribution function $F(x)$ in (1.2), we have

$$E\{g(X_{L(r)}, X_{L(s)})\} = E\{u(X_{L(r-1)}, X_{L(s-1)})\} - E\{u(X_{L(r)}, X_{L(s-1)})\}$$

$$\text{Where } u^*(x, y) = \frac{\partial}{\partial x} u(x, y) = g(x, y), v^*(x)$$

$$\text{And } v^*(x) = \left| \frac{\partial}{\partial x} v(x) \right|$$

Proof: Using (1.3) in (2.2), we have

$$E\{g(X_{L(r)}, X_{L(s)})\} = C_{r,s} \iint_{0 < y < x < \infty} u^*(x, y) [H(x)]^{r-1} [H(y) - H(x)]^{s-r-1} f(y) dx dy$$

Integrating RHS w.r.t “x”, we get as :

$$= C_{r-1, s-1} \iint_{0 < y < x < \infty} u(x, y) [H(x)]^{r-2} [H(y) - H(x)]^{s-r-1} h(x) f(y) dx dy$$

$$- C_{r, s-1} \iint_{0 < y < x < \infty} u(x, y) [H(x)]^{r-1} [H(y) - H(x)]^{s-r-2} h(x) f(y) dx dy$$

and

$$E\{g(X_{L(r)}, X_{L(s)})\} = E\{u(X_{L(r-1)}, X_{L(s-1)})\} - E\{u(X_{L(r)}, X_{L(s-1)})\}$$

Hence the Theorem

Theorem 2.2:

For the distribution function $F(x)$ in (1.2) the recurrence relation for the product Moments of Inverse Weibull distribution is given by:

$$u_{(r),(s)}^{J,K} = \frac{m}{\theta(J-m)} [u_{(r-1),(s-1)}^{(J-m),K} - u_{(r),(s-1)}^{(J-m),K}]$$

Proof: Now $v(x) = \frac{1}{\theta x^m}$ and $g(x,y) = x^J y^K$ This gives

$u(x,y) = \frac{m}{\theta(J-m)} x^{J-m} y^K$, putting in Theorem (2.1), we get the required recurrence relation.

Note, For the ratio, let $K = \mathbf{j}$, then $u_{(r),(s)}^{J,-J} = E\left(\frac{X_{(r)}}{X_{(s)}}\right)^J \quad \forall J$

3. Single Moments

The lower record values are represented by $X_{L(1)}, X_{L(2)}, \dots, X_{L(N)}$. The pdf of $X_{L(n)}$ ($n \geq 2$) is

$$f_{(n)}(x) = \frac{[H(x)]^{n-1}}{(n-1)!} f(x) \quad (3.1)$$

where $H(x) = -\ln F(x)$ $0 < F(x) < 1$

$$h(x) = -\frac{d}{dx} H(x)$$

If g is a Borel measurable function from \mathbb{R}^2 to \mathbb{R} , then

$$E\{g(X_{L(n)})\} = C_n \int_{0 < x < \infty} g(x) [H(x)]^{n-1} f(x) dx \quad (3.2)$$

$0 < x < \infty$

where $C_n = \frac{1}{(n-1)!}$

Theorem 3.1:

For the distribution function $F(x)$ in (1.2), we have

$$E\{g(X_{L(n)})\} = E\{u(x_{L(n-1)})\} - E\{u(x_{L(n)})\}$$

Proof: Using (1.3) in (3.2), we have

$$E\{g(X_{L(n)})\} = C_n \int_{0 < x < \infty} u^{\bullet}(x)[H(x)]^{n-1} F(x) dx$$

Integrating RHS, we get

$$= C_{n-1} \int_{0 < x < \infty} u(x)[H(x)]^{n-2} f(x) dx \quad \blacksquare C_n$$

$$\int_{0 < x < \infty} u(x)[H(x)]^{n-1} f(x) dx$$

and

$$E\{g(X_{L(n)})\} = E\{u(x_{L(n-1)})\} - E\{u(x_{L(n)})\}$$

Hence the Theorem

Theorem 3.2:

For the distribution function $F(x)$ in (1.2) the recurrence relation for the single moments of Inverse Weibull Distribution is given by

$$u_{(n)}^J = \frac{m}{\theta(J-m)} \left[u_{(n-1)}^{J-m} - u_{(n)}^{J-m} \right]$$

Proof: Now $v(x) = \frac{1}{\theta x^m}$ and $g(X_{L(n)}) = x^J$ this gives

$u(X_{L(n)}) = \frac{m}{\theta(J-m)} x^{J-m}$, putting in (3.1) we get the recurrence relation.

Remark: This recurrence relation in Theorem (2.2) between the moments of ratio of two lower record values, quasi-ranges, joint moment generating function, characteristic functions,

whenever they exist can be obtained by setting respectively $g(x,y)$ equal to

$$(x^J y^{-K}), (y-x), e^{T(X+Y)}, e^{iT(X+Y)}.$$

Note: All the above results goes for single lower record values if one replace the function $g(.,.)$ by a function of single variable $g(.)$.

References:

- 1) Ahsanullah, M. & Novzorov, V. B. (2001), "Ordered Random Variables" Nova Science Publishers, Inc. New York
- 2) Ahsanullah, M. (1995). Record Statistics. Nova Science Publishers, Inc. New York.
- 3) Aleem, M. (2004), Some Distributional Properties, Characterization, Ratio, Product and Inverse Moments Of Lower Record values of Inverse Weibull distribution. JOPAS, vol 22(2), 49-56.
- 4) Aleem, M and Pasha G.R. (2003). Ratio, Product and Single Moments Of Order Statistics from Inverse Weibull Distribution. JS, vol x (1), 7-8.
- 5) Gharraph, M.K. (1993). Comparison of Estimators of Location Measures of an Inverse Rayleigh Distribution. The Egyptian Statistical Journal. 37,295-309.
- 6) Mohsin, M (2001). Some Distributional Properties of Lower Record Statistics for inverse Rayleigh Distribution. (Unpublished M. Phil Theses) University of Lahore, Pak. Oct.2001.
- 7) Mukarjee, S. P. and Maiti, s.s.(1996). A Percentile Estimator of the Inverse Rayleigh parameter. IAPR Transactions, 21, 63-65.

- 8) Mukarjee, S. P. and Saren L.K (1984). Bivariate Inverse Rayleigh Distribution in Reliability Studies, Journal of Indian Statistical Association. 22,23-31.
- 9) Voda, V. Gh.(1972). On the “Inverse Rayleigh” Distributed Random Variables, Report in Statistical Applied Research JUSE, 19, 13-21.